# In-Sample and Out-of-Sample Sharpe Ratios in Linear Models

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# I M P E R I A L 🌒 QRT

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  - Findings: We give analytic expressions for the expected in-sample and out-of-sample Sharpe ratios, and find:
    - 1. Higher model complexity inflates in-sample Sharpe ratios
    - 2. Low true Sharpe ratios are vulnerable to overoptimism
    - 3. Short backtests are insufficient to avoid overfitting

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i.e. Past performance does not guarantee future results.



Excluding transaction costs and operational issues, two main problems:

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- Assume the Researcher *fits* and *tests* a linear model on a historical period of length T<sub>1</sub> yielding the in-sample Sharpe ratio SR<sub>IS</sub>.
- What will the expected out-of-sample Sharpe ratio SR<sub>OOS</sub> be on a future period of length T<sub>2</sub>?



Assume a linear prediction model

If  $\beta$  was known, then one could form the portfolio  $\mathbf{w}_t = \mathbf{\Sigma}_{\epsilon}^{-1} \beta \mathbf{s}_t$ , which would have Sharpe ratio

$$\mathsf{SR} = \frac{\mathbb{E}[\mathbf{w}_t^\mathsf{T}\mathbf{r}_{t+1}]}{\sqrt{\mathbb{V}[\mathbf{w}_t^\mathsf{T}\mathbf{r}_{t+1}]}} = \frac{\mathbb{E}\left[(\boldsymbol{\Sigma}_{\epsilon}^{-1}\boldsymbol{\beta}\mathbf{s}_t)^\mathsf{T}(\boldsymbol{\beta}\mathbf{s}_t + \boldsymbol{\epsilon}_{t+1})\right]}{\sqrt{\mathbb{V}\left[(\boldsymbol{\Sigma}_{\epsilon}^{-1}\boldsymbol{\beta}\mathbf{s}_t)^\mathsf{T}(\boldsymbol{\beta}\mathbf{s}_t + \boldsymbol{\epsilon}_{t+1})\right]}}.$$

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Let  $\mathbf{s}_t \sim \mathcal{N}(\mathbf{0}, \mathbf{I}_p)$  and  $\boldsymbol{\epsilon}_{t+1} \sim \mathcal{N}(\mathbf{0}, \boldsymbol{\Sigma}_{\epsilon})$  be mutually independent and IID sequences. Then the Sharpe ratio of the strategy is

$$\mathsf{SR} = rac{\mathrm{tr}(\mathbf{\Gamma})}{\sqrt{2\,\mathrm{tr}(\mathbf{\Gamma}^2)+\mathrm{tr}(\mathbf{\Gamma})}},$$

where

$$\boldsymbol{\Gamma} \coloneqq \boldsymbol{\beta}^{\mathsf{T}} \boldsymbol{\Sigma}_{\epsilon}^{-1} \boldsymbol{\beta}.$$

Thus the Sharpe ratio is increasing in  ${\rm tr}(\Gamma)$  and the ratio  ${\rm tr}(\Gamma^2)/\,{\rm tr}(\Gamma).$ 



In reality, we need to estimate  $\beta$ . If we observe  $T_1$  samples of  $\mathbf{r}_{t+1}$  and  $\mathbf{s}_t$  then by OLS

Stacked returns 
$$\in \mathbb{R}^{m \times T_1}$$
  
 $\widehat{\boldsymbol{\beta}} = \mathbf{R} \mathbf{S}^{\mathsf{T}} (\mathbf{SS}^{\mathsf{T}})^{-1} = \boldsymbol{\beta} + \mathbf{E} \mathbf{S}^{\mathsf{T}} (\mathbf{SS}^{\mathsf{T}})^{-1}$   
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And as  $\mathbb{E} [\mathbf{E}] = \mathbf{0} \implies \mathbb{E} [\widehat{\boldsymbol{\beta}}] = \boldsymbol{\beta}$ . Great!  
So one can then form the portfolio  $\widehat{\mathbf{w}}_t = \boldsymbol{\Sigma}_{\epsilon}^{-1} \widehat{\boldsymbol{\beta}} \mathbf{s}_t$  and earn  
 $\widehat{\mathsf{PnL}}_t = \widehat{\mathbf{w}}_t^{\mathsf{T}} \mathbf{r}_{t+1}$  per time step.

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We can then compute the sample means and variances

Historical Period

$$\begin{split} &\widehat{\mathbb{E}}\left[\left(\widehat{\mathsf{PnL}}_{t}\right)_{t\in\mathcal{T}_{1}}\right] := \frac{1}{T_{1}}\sum_{t=0}^{T_{1}-1}\widehat{\mathsf{PnL}}_{t},\\ &\widehat{\mathbb{V}}\left[\left(\widehat{\mathsf{PnL}}_{t}\right)_{t\in\mathcal{T}_{1}}\right] := \frac{1}{T_{1}-1}\sum_{t=0}^{T_{1}-1}\left(\widehat{\mathsf{PnL}}_{t}-\widehat{\mathbb{E}}\left[\left(\widehat{\mathsf{PnL}}_{t}\right)_{t\in\mathcal{T}_{1}}\right]\right)^{2}, \end{split}$$

and these can be used then to estimate the in-sample Sharpe ratio (and similarly the out-of-sample Sharpe ratio)

$$\mathsf{SR}_{\mathsf{IS}} = \frac{\widehat{\mathbb{E}}\left[(\widehat{\mathsf{PnL}}_t)_{t\in\mathcal{T}_1}\right]}{\sqrt{\widehat{\mathbb{V}}\left[(\widehat{\mathsf{PnL}}_t)_{t\in\mathcal{T}_1}\right]}}, \quad \mathsf{SR}_{\mathsf{OOS}} = \frac{\widehat{\mathbb{E}}\left[(\widehat{\mathsf{PnL}}_u)_{u\in\mathcal{T}_2}\right]}{\sqrt{\widehat{\mathbb{V}}\left[(\widehat{\mathsf{PnL}}_u)_{u\in\mathcal{T}_2}\right]}}.$$
Future Period

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Expanding out the sample P&L,

$$\begin{split} \widehat{\mathsf{PnL}}_{t} &= (\boldsymbol{\Sigma}_{\epsilon}^{-1} \widehat{\boldsymbol{\beta}} \mathbf{s}_{t})^{\mathsf{T}} (\boldsymbol{\beta} \mathbf{s}_{t} + \boldsymbol{\epsilon}_{t+1}) \\ &= (\boldsymbol{\Sigma}_{\epsilon}^{-1} \boldsymbol{\beta} \mathbf{s}_{t})^{\mathsf{T}} \boldsymbol{\beta} \mathbf{s}_{t} + (\boldsymbol{\Sigma}_{\epsilon}^{-1} \boldsymbol{\beta} \mathbf{s}_{t})^{\mathsf{T}} \boldsymbol{\epsilon}_{t+1} \leftarrow \mathsf{Truth} \\ &+ \left( \boldsymbol{\Sigma}_{\epsilon}^{-1} \mathbf{E} \mathbf{S}^{\mathsf{T}} \left( \mathbf{S} \mathbf{S}^{\mathsf{T}} \right)^{-1} \mathbf{s}_{t} \right)^{\mathsf{T}} \boldsymbol{\beta} \mathbf{s}_{t} \leftarrow \mathsf{Misestimation} \\ &+ \left( \boldsymbol{\Sigma}_{\epsilon}^{-1} \mathbf{E} \mathbf{S}^{\mathsf{T}} \left( \mathbf{S} \mathbf{S}^{\mathsf{T}} \right)^{-1} \mathbf{s}_{t} \right)^{\mathsf{T}} \boldsymbol{\epsilon}_{t+1} \leftarrow \mathsf{Overfitting} \end{split}$$

Because when we look at the in-sample P&L,

$$\mathbf{E} = \left( \dots \boldsymbol{\epsilon}_{t+1} \dots \right) \implies \mathbf{E} \not \perp \boldsymbol{\epsilon}_{t+1}.$$



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But, if we look at the out-of-sample P&L, we will have that  $\mathbf{E}$  and  $\mathbf{S}$  are independent of the new realisations of the signals and the residuals.

### $\mathbf{S} \perp\!\!\!\perp \mathbf{s}_u \quad \mathbf{E} \perp\!\!\!\perp \boldsymbol{\epsilon}_{u+1} \iff u \notin \mathcal{T}_1$

# $\implies$ Out-of-Sample Overfitting Disappointment

Q: How large is this difference?



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### Proposition

Given the previous setup, the expected in-sample and out-of-sample mean are,

$$\begin{split} & \mathbb{E}\left[\widehat{\mathbb{E}}\left[\left(\widehat{\mathsf{PnL}}_{t}\right)_{t\in\mathcal{T}_{1}}\right]\right] = \mathrm{tr}(\mathbf{\Gamma}) + \frac{pm}{T_{1}}, \quad \mathbb{E}\left[\widehat{\mathbb{E}}\left[\left(\widehat{\mathsf{PnL}}_{u}\right)_{u\in\mathcal{T}_{2}}\right]\right] = \mathrm{tr}(\mathbf{\Gamma}), \\ & \text{and the expected in-sample and out-of-sample variance are,} \\ & \mathbb{E}\left[\widehat{\mathbb{V}}\left[\left(\widehat{\mathsf{PnL}}_{t}\right)_{t\in\mathcal{T}_{1}}\right]\right] \approx 2\operatorname{tr}(\mathbf{\Gamma}^{2}) + (c_{1}+\widetilde{c}_{1})\operatorname{tr}(\mathbf{\Gamma}) + c_{2} + \widetilde{c}_{2}, \\ & \mathbb{E}\left[\widehat{\mathbb{V}}\left[\left(\widehat{\mathsf{PnL}}_{u}\right)_{u\in\mathcal{T}_{2}}\right]\right] = 2\operatorname{tr}(\mathbf{\Gamma}^{2}) + c_{1}\operatorname{tr}(\mathbf{\Gamma}) + c_{2}, \\ & \text{where } c_{i} \text{ are some variables which increase with } m, p \text{ and decrease} \end{split}$$

with  $T_1$ .

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We propose approximations to the expected in-sample Sharpe ratio  $\mathbb{E}[SR_{IS}]$  and the expected out-of-sample Sharpe ratio  $\mathbb{E}[SR_{OOS}]$ 

$$\begin{split} \mathbb{E}[\mathsf{SR}_{\mathsf{IS}}] &\approx \mathsf{SR}_{\mathsf{EIS}} = \frac{\mathbb{E}\left[\widehat{\mathbb{E}}\left[(\widehat{\mathsf{PnL}}_t)_{t\in\mathcal{T}_1}\right]\right]}{\sqrt{\mathbb{E}\left[\widehat{\mathbb{V}}\left[(\widehat{\mathsf{PnL}}_t)_{t\in\mathcal{T}_1}\right]\right]}},\\ \mathbb{E}[\mathsf{SR}_{\mathsf{OOS}}] &\approx \mathsf{SR}_{\mathsf{EOOS}} = \frac{\mathbb{E}\left[\widehat{\mathbb{E}}\left[(\widehat{\mathsf{PnL}}_u)_{u\in\mathcal{T}_2}\right]\right]}{\sqrt{\mathbb{E}\left[\widehat{\mathbb{V}}\left[(\widehat{\mathsf{PnL}}_u)_{u\in\mathcal{T}_2}\right]\right]}}, \end{split}$$

and we use these to study the replication ratio  $\frac{SR_{EOOS}}{SR_{EIS}}$  .

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How is the replication ratio affected by increasing the number of assets m and signals p?

Example: 10 year backtest (2520 days) with in-sample SR 2.



(a)

- If we don't have a predictive model and instead just estimate the drift and covariance of our assets, what is the replication ratio? i.e. statically hold the portfolio  $\mathbf{w} = \Sigma^{-1} \mu$ .
- Kan, Wang, and Zheng [KWZ22]: let the true SR be θ = <sup>w<sup>T</sup>μ</sup>/<sub>√w<sup>T</sup>Σw</sub> = √μ<sup>T</sup>Σ<sup>-1</sup>μ, the authors compute the expected in-sample SR E[θ] and out-of-sample SR E[θ].



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Similar haircuts when p=1, but not exactly the same as dynamic vs static.

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(a)

How does this work when our assumptions are violated?

- Using 12 commodity futures from 1998 to 2023 compute rolling 5-day, 1-year and 5-year t-statistics as signals
- Fit an AR(1) model with t-distributed or Normally distributed residuals:

$$\mathbf{r}_{t+1} = \boldsymbol{\beta} \mathbf{s}_t + \boldsymbol{\epsilon}_{t+1},$$
$$\mathbf{s}_t = \boldsymbol{\Phi} \mathbf{s}_{t-1} + \mathbf{u}_t,$$

 Simulate new samples using this model to check impact of assumption violations



#### Commodity Futures



Replication Ratio by IID/AR Signals and Normal/t residuals

Figure: Distribution of the replication ratio, with iid or AR(1) signals and Normal or t distributed residuals.

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#### Commodity Futures



Figure: Distribution of the replication ratio by out-of-sample Sharpe ratio.

Peter Muller (Founder, PDT Partners) [Mul01]:

In my opinion it is far better to refine an individual strategy...than to attempt to put together lots of weaker strategies...I would much rather have a single strategy with an expected Sharpe ratio of 2 than a strategy that has an expected Sharpe ratio of 2.5 formed by putting together five supposedly uncorrelated strategies each with an expected Sharpe ratio of 1.

Nick Patterson (RenTech) [Pat16]:

It's funny that I think the most important thing to do in data analysis is to do the simple things right. So, here's a kind of non-secret about what we did at Renaissance: in my opinion, our most important statistical tool was simple regression with one target and one independent variable.



Common sense prevails:

- $1. \ \mbox{Use the longest backtest you can}$
- 2. Don't use too many signals

3. Don't trust low Sharpe ratios (or too high Sharpe ratios!) Good luck!



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